Lecture 11 | 07.05.2024

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Regression models beyond typical data

Linear regression models and beyond

- □ Linear models... but **the truth is (almost) never linear!** (the linearity property is used as a good and easy approximation)
- □ Nevertheless, it is convenient to have simple assumptions... (but there are many different issues that can go wrong...)
- Recall, that there are a few levels of linearity in the model (linearity of the predictor, linearity of the expectation, linearity of LS)

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the data are too flexibile (higher order approximations/splines)
the data are too irregular (piecewise approximation)
the data are too complex (additive models)
the data are too volatile (robust estimation approaches)
the nature of Y contradicts the linear model (GLM)
and many more reasons (and way more alternatives)

Recap: Linear regression framework

- □ for a generic random vector $(Y, \mathbf{X}^{\top})^{\top} \in \mathbb{R}^{p+1}$ we assume an unknown population model $Y = \mathbf{X}^{\top} \boldsymbol{\beta} + \varepsilon$ for an unknown vector $\boldsymbol{\beta} \in \mathbb{R}^{p}$
- □ for a random sample $\{(Y_i, \mathbf{X}_i^{\top})^{\top}; i = 1, ..., n\}$ drawn from the joint distribution $F_{(Y, \mathbf{X})}$ we have data model $Y_i = \mathbf{X}_i^{\top} \boldsymbol{\beta} + \varepsilon_i$
- □ the data model can be also expressed as $\boldsymbol{Y} | \mathbb{X} \sim (\mathbb{X}\beta, \sigma^2 \mathbb{I})$, for the random vector $\boldsymbol{Y} = (Y_1, \ldots, Y_n)^\top$ and $\mathbb{X} = (\boldsymbol{X}_1, \ldots, \boldsymbol{X}_n)^\top \in \mathbb{R}^{n \times p}$, rank $(\mathbb{X}) = p$

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- Moreover:

$$\begin{array}{l} \widehat{\boldsymbol{\beta}} = (\mathbb{X}^{\top}\mathbb{X})^{-1}\mathbb{X}^{\top}\boldsymbol{Y} \text{ and } \widehat{\boldsymbol{Y}} = \mathbb{X}\widehat{\boldsymbol{\beta}} \\ \hline \boldsymbol{Y} = \mathbb{H}\boldsymbol{Y} + \mathbb{M}\boldsymbol{Y}, \text{ where } \mathbb{H} = \mathbb{X}(\mathbb{X}^{\top}\mathbb{X})^{-1}\mathbb{X}^{\top} = (h_{ij})_{i,j=1}^{n} \text{ and } \mathbb{M} = (m_{ij})_{i,j=1}^{n} \\ \hline \boldsymbol{Y} = \mathbb{M}\boldsymbol{Y} = (\mathbb{I} - \mathbb{H})\boldsymbol{Y} = \boldsymbol{Y} - \widehat{\boldsymbol{Y}} = (U_{1}, \ldots, U_{n})^{\top} \\ \hline SSe = \sum_{i=1}^{n} (Y_{i} - \widehat{Y}_{i})^{2} = \|\boldsymbol{Y} - \widehat{\boldsymbol{Y}}\|_{2}^{2} = \|\boldsymbol{U}\|_{2}^{2} \text{ and } MSe = SSe/(n-p) \\ \hline \text{ standardized residuals } V_{i} = U_{i}/\sqrt{MSe \cdot m_{ii}}, \text{ if } m_{ii} > 0 \end{array}$$

Linear regression models

Least squares and the linear regression models based on the LS minimization are, in general, very sensitive (non-robust) with respect to atypical (non-normal, skewed, and heavy-tailed) data... But it is not straightforward to say what **atypical** actually means...

Two common concepts are:

Outlying observations

an outlying observation in some regression model $Y = \mathbf{X}^{\top} \boldsymbol{\beta} + \varepsilon$ is an observation for which the response expectation $(E[Y|\mathbf{X}])$ does not follow the assumed model $\mathbf{X}^{\top} \boldsymbol{\beta}$, respectively, it is an observation $\iota \in \{1, \ldots, n\}$ such that $E[Y_{\iota}|\mathbb{X}_{\iota}] \neq \mathbf{X}_{\iota}^{\top} \boldsymbol{\beta}$ (i.e., $E[Y_{\iota}|\mathbb{X}_{\iota}] = \mathbf{X}_{\iota}^{\top} \boldsymbol{\beta} + \gamma$)

Leverage points

a leverage point in some regression model $Y = \mathbf{X}^{\top} \boldsymbol{\beta} + \boldsymbol{\varepsilon}$ is an observation which is, in some sense, unusual with respect to the regresor values in \mathbf{X} .

Outlying observations and leverage points

- □ It is a well-known fact that a few bad leverage points or outlyiers can result in a (very) poor fit to the bulk of the data
- Morever, this can be even the case when using more robust alternatives that should avoid this drawback
- outlying observations and leverage points are of different nature—either of them can appear in the data (model) but they can also appear simultanously
- □ different strategies are proposed in the literature to deal with the outliers, with the leverage points, or both of them simultaneously

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- □ different strategies are proposed in the literature to deal with the outliers, with the leverage points, or both of them simultaneously
- for a simple illustration, consider a problem of a simple mean and a simple median calculated from some univariate random sample... while the average is sensitive with respect to just one outlying observation, the sample median is way more robust...

Motivation

Outlying observations: mathematically

□ for a regression (data) model $Y_i = \mathbf{X}_i^\top \beta + \varepsilon_i$ and some observation $\iota \in \{1, \ldots, n\}$ (fixed) we define the following two models:

Leave-one-out model

$$\mathcal{M}_{-\iota}: \quad \mathbf{Y}_{-\iota} | \mathbb{X}_{-\iota} \sim (\mathbb{X}_{-\iota} \boldsymbol{\beta}, \sigma^2 \mathbb{I}_{n-1})$$

where $-\iota$ denotes the observation which is omitted

Outlyier model

$$\mathcal{M}_{\iota}: \quad \mathbf{Y}_{\iota} | \mathbb{X}_{\iota} \sim (\mathbb{X}_{\iota} \boldsymbol{\beta} + \mathbf{j}_{\iota}^{\top} \gamma_{\iota}, \sigma^{2} \mathbb{I}_{n-1})$$

where ι denotes the observation which is outlying and j_{ι} is a unit vector with one on the position $\iota \in \{1, \ldots, n\}$

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□ It can be proved, that the residual sum of squares in both models are the same (meaning that $SSe_{-\iota} = SSe_{\iota}$). The vector $\hat{\beta}_{-\iota}$ solves the normal equations in the model $\mathcal{M}_{-\iota}$ if and only if $(\hat{\beta}_{\iota}^{\top}, \hat{\gamma})_{\iota}^{\top}$ solves the normal equations in \mathcal{M}_{ι} , where $\hat{\beta}_{-\iota} = \hat{\beta}_{\iota}$ and $\hat{\gamma} = Y_{\iota} - X_{\iota}^{\top} \hat{\beta}_{-\iota}$

Detection of outlying observations

- □ for any $\iota \in \{1, ..., n\}$ we denote $\widehat{Y}_{[\iota]} = \mathbf{X}_{\iota}^{\top} \widehat{\beta}_{-\iota}$ which is acually a least squares estimate of $\mu_{\iota} = E[Y_{\iota} | \mathbf{X}_{\iota}]$ but using only n 1 observations for $i = 1, ..., \iota 1, \iota + 1, ..., n$
- □ the whole vector $\widehat{\mathbf{Y}}$ can be estimated by using a leave-one-out model, obtaining $\widehat{\mathbf{Y}}_{[]} = (\widehat{Y}_{[1]}, \dots, \widehat{Y}_{[n]})$
- It also holds that

$$\widehat{\boldsymbol{\gamma}}_{\iota} = \widehat{\boldsymbol{\gamma}}_{\iota} - \boldsymbol{X}_{\iota}^{\top} \widehat{\boldsymbol{\beta}}_{-\iota} = \boldsymbol{Y}_{\iota} - \widehat{\boldsymbol{Y}}_{[\iota]} = \frac{\boldsymbol{U}_{\iota}}{m_{\iota\iota}} \widehat{\boldsymbol{\beta}}_{-\iota} = \widehat{\boldsymbol{\beta}}_{\iota} = \widehat{\boldsymbol{\beta}} - \frac{\boldsymbol{U}_{\iota}}{m_{\iota\iota}} (\mathbb{X}^{\top} \mathbb{X})^{-1} \boldsymbol{X}_{\iota} \widehat{\boldsymbol{SSe}}_{-\iota} = SSe_{\iota} = SSe - \frac{\boldsymbol{U}_{\iota}^{2}}{m_{\iota\iota}} = SSe - MSe(V_{\iota}^{2})$$

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$$\begin{array}{l} \square \ \widehat{\gamma}_{\iota} = \widehat{Y}_{\iota} - \boldsymbol{X}_{\iota}^{\top} \widehat{\beta}_{-\iota} = Y_{\iota} - \widehat{Y}_{[\iota]} = \frac{U_{\iota}}{m_{\iota\iota}} \\ \square \ \widehat{\beta}_{-\iota} = \widehat{\beta}_{\iota} = \widehat{\beta} - \frac{U_{\iota}}{m_{\iota\iota}} (\mathbb{X}^{\top} \mathbb{X})^{-1} \boldsymbol{X}_{\iota} \\ \square \ SSe_{-\iota} = SSe_{\iota} = SSe - \frac{U_{\iota}^{2}}{m_{\iota\iota}} = SSe - MSe(V_{\iota}^{2}) \end{array}$$

- □ thus, the original regression model $\boldsymbol{Y} | \mathbb{X} \sim (\mathbb{X}\beta, \sigma^2 \mathbb{I})$ can be used o detect outlying observations in the model
- □ from the inferential point of view, it is also easy to test the null hypothesis H_0 : $\gamma_{\iota} = 0$ (detection of an outlier)

Something to keep in mind

- Two or more outliers next to each other can hide each other
- A notion of outlier is always relative to considered model—an observation which is an outlier with respect to one model is not necessarily an outlier with respect another model
- Outlier can also suggest that a particular observation is a data-error that must be corrected
- If some observation is indicated to be an outlier, it should always be explored
- □ Often, identification of outliers with respect to some model is of primary interest (e.g., credit card transactions)

Cross-validation (CV)

- Cross-validation is a very popular and commonly used statistical techniques (also applied in regression) which is based on the vector $\widehat{\mathbf{Y}}_{[1]} = (\widehat{Y}_{[1]}, \dots, \widehat{Y}_{[n]})^{\top}$ (so-called leave-one-out CV)
- \Box the residual $U_{\iota} = Y_{\iota} \widehat{Y}_{\iota}$ for some observation $\iota \in \{1, \ldots, n\}$ may be considered to be too optimistic, because the value of Y_{ι} was used to train the model—i.e., to estimate β and to obtain $\widehat{\mathbf{Y}} = \mathbb{X}\widehat{\beta}$ (and also \widehat{Y}_{ι})
- slightly less optimistic residual (sometimes also called the deleted residual) obtained by the quantity $\widehat{\gamma}_{\iota} = Y_{\iota} - \widehat{Y}_{[\iota]} = U_{\iota}/m_{\iota\iota}$ because the value of Y_{ι} is not estimated by using the data that does not contain Y_{i} itself
- \Box more general concepts (so-called *k*-fold cross-validation) are also know in the literature and these techniques are commonly used in regression modelling in practice

Leverage points

- **Q** considering the hat matrix $\mathbb{H} = \mathbb{X}(\mathbb{X}^{\top}\mathbb{X}^{-1})\mathbb{X}^{\top} = (h_{ii})_{i,i=1}^{n}$, the element h_{ii} for some $i \in \{1, \ldots, n\}$ is called the leverage of X_i
- \square it is easy to show that $\sum_{i=1}^{n} h_{ii} = tr(\mathbb{H}) = tr(\mathbb{Q}\mathbb{Q}^{\top}) = tr(\mathbb{Q}^{\top}\mathbb{Q}) = p$ thus, the average leverage is $\overline{h} = \frac{1}{n} \sum_{i=1}^{n} h_{ii} = k/n$
- some rule-of-thumb for identifying leverage points uses the criterion $h_{ii} > 3k/n$
- Other alternatives include
 - DEBETAS

the analysis of the effect of a particular observation on the estimates of some parameter β_i

DFFITS

the analysis of the effect of the ι^{th} observation on the estimates of Y_{ι}

□ COVRATIO

the analysis of the effect of a particular observation on the estimates of the parameter vector β

Cook distance

the analysis of the effect of a particular observation on the estimates of the mean vector $\boldsymbol{\mu} = E[\boldsymbol{Y}|\mathbb{X}]$

How to deal with outliers and leverage points

Different techniques and methodological approaches can be used to deal with the outlying observations, with the leverage points, for both simultaneously...

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Different techniques and methodological approaches can be used to deal with the outlying observations, with the leverage points, for both simultaneously...

- naive methods use the principle of deleting bad outiers and bad leverage poitns... this should, however, never be done automatically—a proper exploratory is needed
- more advanced methods used (iterative) re-weighted least squares where the weights are determed by some of the criterion mentioned above
- robust regression alternative which are not that much sensitive to the ouliers, leverage points, or both simultaneously can be used instead (e.g., the median regression)

Motivation

Summary

Outlying observations

- unusual observations with respect to the observed values of the response
- outliers may have serious consequences with respect to the final fit
- □ different recommendations are used to detect and classify outliers
- various alternatives are proposed to incorporate outliers into the model

Leverage points

- unusual observations with respect to the values of the covariates
- leverage points may also have serious impact on the final fit
- □ different tools are used to explore leverage points
- D modifications of the regression framework are used to bad leverage points